

On the interplay between distortion, mean value and Haezendonck-Goovaerts risk measures

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Mean value risk measures

- Definition: continuity condition

- ▶ Let X_{aq} be a Bernoulli risk with

$$\Pr(X_{aq} = a) = q$$

$$\Pr(X_{aq} = 0) = 1 - q.$$

For a fixed $a > 0$, the risk measure H satisfies the *continuity condition* if, and only if, $H(X_{aq})$ is strictly increasing for $0 \leq q \leq 1$, with $H(X_{a0}) = 0$ and $H(X_{a1}) = a$.

- Definition: mean value risk measure

- ▶ A risk measure H is said to be generated by the *mean value principle* if there exists a strictly increasing function v such that $v(H(X)) = E[v(X)]$.

- Theorem:

- ▶ A risk measure H satisfying the continuity condition is iterative if, and only if, it is generated by the mean value principle.

Comparing risk measures

- Definition: comparability

- ▶ Two mean value principles H , with strictly increasing v_1 and v_2 , are comparable in case for all bounded risks X

$$H(X, v_1) \leq H(X, v_2),$$

or the reverse inequality, with $H(X, v) = v^{-1}(\mathbb{E}[v(X)])$.

- Theorem:

- ▶ Let v_1 and v_2 be two continuous and strictly increasing functions in \mathbb{R} . A necessary and sufficient condition for $H(X, v_1)$ and $H(X, v_2)$ to be comparable is that the function

$$h = v_2 v_1^{-1}$$

satisfies

$$h(\mathbb{E}[X]) \leq \mathbb{E}[h(X)],$$

or the reverse inequality, for all risks $X \in \mathcal{B}$. Hence, h has to be a convex or a concave function.

Risk Ordering

- Compare the "risk premiums" $H(X)$ and $H(Y)$.
 - ▶ If v is increasing we have for $X \leq_1 Y$ that $E[v(X)] \leq E[v(Y)]$, and as a consequence that $H(X) \leq H(Y)$.
 - ▶ If $E[(X - t)_+] \leq E[(Y - t)_+]$ for all t , then $E[v(X)] \leq E[v(Y)]$ for all non-decreasing convex functions v , and consequently $H(X) \leq H(Y)$
- Convex order:

$$E[(X - t)_+] \leq E[(Y - t)_+] \text{ for all } t, \\ \text{and } E[X] = E[Y].$$

These conditions are equivalent to

$$E[(t - X)_+] \leq E[(t - Y)_+] \text{ for all } t, \\ \text{and } E[X] = E[Y].$$

Optimal risk measures

- Premium calculation:

- ▶ Consider an exponential function v .
- ▶ Exponential premium:

$$H(X) = \frac{1}{\alpha} \ln \mathbb{E} \left[e^{\alpha X} \right].$$

- Consider the following inequalities:

$$\begin{aligned} \frac{1}{\alpha} \ln \mathbb{E} \left[e^{\alpha X} \right] &= \frac{1}{\alpha} \int_0^\alpha \frac{\mathbb{E} \left[e^{sX} X \right]}{\mathbb{E} \left[e^{sX} \right]} ds \\ &\leq \frac{\mathbb{E} \left[e^{\alpha X} X \right]}{\mathbb{E} \left[e^{\alpha X} \right]}, \end{aligned}$$

- ▶ Esscher premium.

Optimal risk measures

- Weighted Esscher premiums:

$$H(X) = \int_{-\infty}^{+\infty} \frac{\mathbb{E}[e^{tX} X]}{\mathbb{E}[e^{tX}]} dG(t),$$

where $G : \mathbb{R} \rightarrow [0, 1]$ and G is concave on $(0, +\infty)$ and convex on $(-\infty, 0)$.

- Consequently $H(X)$ can be written as $\mathbb{E}^*[X]$, using the differential

$$dF_X^{G(\cdot)}(x) = \int_{-\infty}^{+\infty} \frac{e^{tX} dG(t)}{\mathbb{E}[e^{tX}]} dF_X(x). \quad (1)$$

Application of the mean value principle to generate distortion risk measures

- Distortion risk measure:

- ▶ g is increasing with $g(0) = 0$ and $g(1) = 1$. The *distortion risk measure* $\rho_g(X)$ is:

$$\rho_g(X) = \int_0^\infty g(1 - F_X(x)) dx \quad (2)$$

$$= \int_0^1 F_X^{-1}(y) g'(1 - y) dy. \quad (3)$$

- ▶ $g_1(y) \geq g_2(y)$ for all $y \implies \rho_{g_1}(X) \geq \rho_{g_2}(X)$.
- ▶ $g_2(x) = x$ implies that $\rho_g(X) \geq E[X]$ for any distortion risk measure ρ_g , with $g(x) \geq x$.

- Comonotonicity:

- ▶ For all (x_1, y_1) and $(x_2, y_2) : x_1 \leq x_2$ and $y_1 \leq y_2$ or the other way around.
- ▶ X^c and Y^c are maximal dependent.
- ▶ Quantiles of a comonotonic sum: $F_{X^c + Y^c}^{-1}(p) = F_{X^c}^{-1}(p) + F_{Y^c}^{-1}(p)$.

Distortion risk measure

- Comonotonic risks:

$$\rho_g(X^c + Y^c) = \rho_g(X^c) + \rho_g(Y^c).$$

- Bernoulli risk X_{aq} :

$$\rho_g(X_{aq}) = (1 - q)0 + \int_{1-q}^1 ag'(1 - y)dy = ag(q).$$

- $\rho_g(X_{aq})$ as a mean value risk measure:

- ▶ mean value function $v : v(\rho_g(X_{aq})) = E[v(X_{aq})]$
- ▶ which is equivalent with: $v(ag(q)) = qv(a)$.
- ▶ It follows that: $g(q) = q$ and $v'(aq) = v'(a)$.
- ▶ Hence: v and g have to be linear functions.

Characterization of Wang's class of premium principles

Some desirable properties for premium principles

- For any two risks X and Y : $1 - F_X(x) \leq 1 - F_Y(x)$ for all $x \geq 0$ implies $H(X) \leq H(Y)$.
- For comonotonic risks: $H(X + Y) = H(X) + H(Y)$.
- $H(1) = 1$.
- $H(X) \geq E[X]$.
- For any $d \geq 0$: $\lim_{d \rightarrow +\infty} H[\min(X, d)] = H(X)$.

Characterization of Wang's class of premium principles

- Lemma:

- ▶ Assume a premium principle H , having the properties 1-4. Then there exists a unique distortion function g such that for all discrete risks X , with only finitely many mass points, we have that:

$$H(X) = \int_0^{+\infty} g(1 - F_X(x)) dx.$$

Furthermore, $g(q) \geq q$ for all $q \in [0, 1]$.

- Theorem:

- ▶ Assume that H satisfies properties 1-5. Then there exists a unique distortion function g , with $g(q) \geq q$ for all $q \in [0, 1]$, such that for all X , we have that:

$$H(X) = \int_0^{+\infty} g(1 - F_X(x)) dx$$

Risk measures for capital requirements

- Risk in the right tail:

- ▶ Use the transformed r.v. Z :

$$Z = \frac{\left(F_X^{-1}(U) - t\right)_+}{\rho - t}.$$

- ▶ $Z > 1$: residual risk.
- ▶ $Z < 1$: residual gains.

- Measuring the risk in the right tail:

- ▶ Mean value risk measure:

$$v(H(Z)) = E[v(Z)].$$

- ▶ take: $v(x) = x$ for $x < 1$ and $v(x) > x$ for $x > 1$.
- ▶ $\rho_l(X)$ is determined via:

$$H(Z) = \int_0^1 \frac{\left(F_X^{-1}(u) - t\right)_+}{\rho_l(X, t) - t} du = 1 - \alpha.$$

Risk measures for capital requirements (cont'd)

- Solving for $\rho_l(X, t)$

- ▶ $\rho_l(X, t)$ can be determined from:

$$E \left[\frac{\left(F_X^{-1}(U) - t \right)_+}{\rho_l(X, t) - t} \right] = 1 - \alpha,$$

or:

$$\begin{aligned} \rho_l(X, t) &= t + \frac{1}{1 - \alpha} \int_t^{+\infty} (x - t)_+ dF_X(x) \\ &= t + \int_t^{+\infty} \frac{1 - F_X(x)}{1 - \alpha} dx. \end{aligned} \quad (4)$$

- $\rho_l(X, t) - t$ is not a distortion risk measure

- ▶ $\rho_l(X, t) - t$ seems to be a distortion risk measure
- ▶ Distortion function $g(x) = \frac{x}{1 - \alpha}$.
- ▶ $g(1) = \frac{1}{1 - \alpha} > 1$: hence g is not a distortion function.

Risk measures for capital requirements (cont'd)

- $\rho_g(X, t) - t$ is a distortion risk measure

- ▶ Take $t = F_X^{-1}(\alpha)$. Define g as:

$$g(x) = \min \left\{ \frac{x}{1-\alpha}, 1 \right\}. \quad (5)$$

- ▶ $\rho_g(X, F_X^{-1}(\alpha))$ is equal to:

$$\begin{aligned} \rho_g(X, F_X^{-1}(\alpha)) &= \int_0^{+\infty} g(1 - F_X(x)) dx \\ &= \int_0^{F_X^{-1}(\alpha)} 1 dx + \int_{F_X^{-1}(\alpha)}^{+\infty} \frac{1 - F_X(x)}{1 - \alpha} dx \\ &= t + \int_t^{+\infty} \frac{1 - F_X(x)}{1 - \alpha} dx. \end{aligned}$$

- Conclusion:

$$\rho_g(X, F_X^{-1}(\alpha)) = \rho_l(X, F_X^{-1}(\alpha)).$$

Risk measures for capital requirements (cont'd)

- comonotonic risks X_1^c and X_2^c

- ▶ Quantiles are additive: $F_{X_1^c + \dots + X_n^c}^{-1}(\beta) = \sum_{i=1}^n F_{X_i^c}^{-1}(\beta)$.
- ▶ Right tail of the comonotonic sum can be decomposed into a sum of random variables:

$$\left(X_1^c + \dots + X_n^c - F_{X_1^c + \dots + X_n^c}^{-1}(\beta) \right)_+ = \sum_{i=1}^n \left(X_i^c - F_{X_i^c}^{-1}(\beta) \right)_+.$$

- For $t = F_X^{-1}(\beta)$, with $\beta < 1$:

$$\begin{aligned} & \rho_l(X_1^c + X_2^c) - F_{X_1^c}^{-1}(\beta) - F_{X_2^c}^{-1}(\beta) \\ &= \rho(X_1) - F_{X_1}^{-1}(\beta) + \rho(X_2) - F_{X_2}^{-1}(\beta), \end{aligned}$$

or

$$\begin{aligned} & \rho_l \left(\left(X_1^c + X_2^c - F_{X_1^c + X_2^c}^{-1}(\beta) \right)_+ \right) \\ &= \rho_l \left(\left(X_1^c - F_{X_1^c}^{-1}(\beta) \right)_+ \right) + \rho_l \left(\left(X_2^c - F_{X_2^c}^{-1}(\beta) \right)_+ \right). \end{aligned}$$

Risk measures for capital requirements (cont'd)

- Denote the residual risk measure, given the capital t , as $\pi(X, t) = \rho - t$. Then $\pi(X^c, t)$ for $X^c = X_1^c + X_2^c + \dots + X_n^c$ is determined as (in case $v(x) = x$):

$$1 - \alpha = \int_{\beta}^1 \frac{\sum_j \left(F_{X_j^c}^{-1}(u) - F_{X_j^c}^{-1}(\beta) \right)_+}{\pi_I(X^c, t)} du,$$

where $\pi_I(X^c, t)$ is obtained by means of a particular mean value principle with a linear function v .

- Conclusion:
 - $\pi_I(X, t)$ is derived out of a mean value principle.
 - $\pi_I(X, t)$ is comonotone additive.
 - When $t = F_X^{-1}(\alpha)$, $\pi_I(X, t)$ is expressed as a distortion risk measure.

Measuring the tail risk

Application of a mean value risk measure

- Introduction:

- ▶ Consider a random variable X and a function φ , strictly increasing with $\varphi(0) = 0$, $\varphi(1) = 1$, $\varphi(+\infty) = +\infty$:

$$\Pr(X > \rho) = \Pr(X - t > \rho - t) \leq \mathbb{E} \left[\varphi \left(\frac{(X - t)_+}{\rho - t} \right) \right].$$

- Definition :

$$\mathbb{E} \left[\varphi \left(\frac{(X - t)_+}{\rho - t} \right) \right] = 1 - \alpha, \quad (6)$$

always has a solution. This solution is denoted by $\rho_\varphi(X, t)$.

Haezendonck-Goovaerts risk measure risk measure

- Definition:

- ▶ Let φ be a strictly increasing function with $\varphi(0) = 0$, $\varphi(1) = 1$, $\varphi(+\infty) = +\infty$ and let $\alpha \in (0, 1)$. The *Haezendonck-Goovaerts risk measure* is denoted by $\rho_\varphi(X)$:

$$\rho_\varphi(X) = \inf_{-\infty < t < \max[X]} \rho_\varphi(X, t),$$

where $\rho_\varphi(X, t)$ is the solution of equation (6).

- Positive Homogeneous and translation invariant:

- ▶ With $t = F_X^{-1}(\beta)$:

$$1 - \alpha = E \left[\varphi \left(\frac{(F_X^{-1}(U) - F_X^{-1}(\beta))_+}{\rho_\varphi(X, t) - F_X^{-1}(\beta)} \right) \right], \quad (7)$$

- ▶ We can see that:

$$\begin{aligned} \rho_\varphi(aX, t) &= a\rho_\varphi(X, t), \text{ for } a > 0, \\ \rho_\varphi(a + X, t) &= a + \rho_\varphi(X, t) \text{ for } a \in \mathbb{R}. \end{aligned}$$

Solving for the Haezendonck-Goovaerts risk measure risk measure

- The Haezendonck-Goovaerts risk measure risk measure $\rho_\varphi(X)$ is determined as the solution of the system of equations:

$$1 - \alpha = \int_t^{+\infty} \varphi\left(\frac{(x-t)_+}{\rho-t}\right) dF_X(x), \quad (8)$$

$$\rho = t + \frac{\int_t^{+\infty} \varphi'\left(\frac{(x-t)_+}{\rho-t}\right) (x-t)_+ dF_X(x)}{\int_t^{+\infty} \varphi'\left(\frac{(x-t)_+}{\rho-t}\right) dF_X(x)}. \quad (9)$$

- Special cases:

$$\blacktriangleright \varphi(x) = x: \quad \rho = F_X^{-1}(\alpha) + \frac{1}{1-\alpha} \mathbb{E}\left[\left(X - F_X^{-1}(\alpha)\right)_+\right].$$

$$\blacktriangleright \varphi(x) = e^{\alpha x}: \quad \rho = t + \frac{\mathbb{E}\left[(X-t)_+ e^{\frac{\alpha X}{\rho-t}}\right]}{\mathbb{E}\left[e^{\frac{\alpha X}{\rho-t}}\right]}.$$

Haezendonck-Goovaerts risk measure risk measure and TVAR

- The risk measure $\rho_l(S, t)$:



$$\rho_l(S, t) = t + \frac{1}{1-\alpha} \mathbb{E}[(S-t)_+], \quad (10)$$

is the solution of the equation:

$$\mathbb{E}\left[\frac{(S-t)_+}{\rho_l(S, t) - t}\right] = 1 - \alpha.$$

- ▶ We have that $\text{TVaR}(S, \alpha) = \rho_l(S, \text{VaR}[S, \alpha])$.

- Comparability of the mean value principles:

- ▶ $\mathbb{E}\left[\varphi\left(\frac{(S-t)_+}{\rho_\varphi(S, t) - t}\right)\right] = 1 - \alpha$, gives

$$\rho_\varphi(S, t) > \rho_l(S, t).$$

- Conclusion:

$$\text{TVaR}(S, \alpha) \leq \rho_\varphi(S, \text{VaR}[S, \alpha]). \quad (11)$$

Haezendonck-Goovaerts risk measure risk measure for two point distributions

- Bernoulli random variable B_q :

- ▶ $\Pr(B_q = 1) = 1 - \Pr(B_q = 0) = q.$
- ▶ Using the function $\varphi(x) = x$: $\rho_l(B_q) = \min\left\{\frac{q}{1-\alpha}, 1\right\}.$
- ▶ For a general choice of $\varphi(x)$: $\rho_\varphi(B_q) = \min\left\{\frac{1}{\varphi^{-1}\left(\frac{1-\alpha}{q}\right)}, 1\right\}.$

- Consider the distribution $(aB_q - t)_+$, $a < t$:

- ▶ $\Pr(aB_q - t = a - t) = 1 - \Pr(aB_q - t = 0) = q.$
- ▶ We get:

$$\rho_l((aB_q - t)_+) = (a - t) \min\left\{\frac{q}{1-\alpha}, 1\right\},$$

$$\rho_\varphi((aB_q - t)_+) = (a - t) \min\left\{\frac{1}{\varphi^{-1}\left(\frac{1-\alpha}{q}\right)}, 1\right\}$$

The connection between the Haezendonck-Goovaerts risk measure risk measure and distortion risk measures

- Theorem:

- ▶ Consider a Bernoulli risk B_q and a function $g(q)$ which is a distortion measure function such that $g(q)$ is increasing for $0 < q < 1 - \alpha$ and $g(q) = 1$ for $1 - \alpha < q \leq 1$. A sufficient condition for the existence of a convex function φ for which the equality $\rho_\varphi(B_q) = \rho_g(B_q)$ holds, is that $g(q)$ is concave for $q \leq 1 - \alpha$.

Proof:

- For a Bernoulli risk B_q , the equality $\rho_\varphi(B_q) = \rho_g(B_q)$ gives:

$$\varphi\left(\frac{1}{g(q)}\right) = \frac{1-\alpha}{q}.$$

- Let $c(q) = \frac{1}{g(q)}$, then

$$c'(q) = -\frac{g'(q)}{g^2(q)} \leq 0,$$

$$c''(q) = -\frac{g''(q)g(q) - 2(g'(q))^2}{g^3(q)} \geq 0,$$

for a concave distortion function g .

Proof (Cont'd):

- Hence, $c(q)$ is a decreasing convex function for $0 < q < 1 - \alpha$.
- If we set $x = c(q)$ (or $q = c^{-1}(x)$), we have that

$$\frac{dx}{dq} = c'(q),$$

- which also means that

$$\begin{aligned}\frac{dq}{dx} &= \frac{1}{c'(q)} \\ &= \frac{1}{c'(c^{-1}(x))} \\ &< 0.\end{aligned}$$

Proof (Cont'd):

- For the second derivative we find

$$\frac{d^2 q}{dx^2} = -\frac{c''(q)}{(c'(q))^3} \geq 0.$$

- Consequently, $c^{-1}(x)$ is a decreasing and convex function. The function $\varphi(x)$ can be expressed in terms of $c^{-1}(x)$ as follows:

$$\varphi(x) = \frac{1 - \alpha}{c^{-1}(x)},$$

and

$$\varphi'(x) = -(1 - \alpha) \frac{(c^{-1}(x))'}{(c^{-1}(x))^2}.$$

- The second derivative can be written as:

$$\varphi''(x) = -(1 - \alpha) \frac{(c^{-1}(x))'' c^{-1}(x) - 2(c^{-1}(x))'^2}{(c^{-1}(x))^3}.$$

Proof (Cont'd):

- Using the function $c(q)$, we can write

$$\varphi''(x) = \frac{-(1-\alpha)}{q^3} \left[-\frac{c''(q)}{c'(q)^3} q - \frac{2}{c'(q)^2} \right]$$

- and directly in terms of the distortion function as

$$\varphi''(x) = \frac{-(1-\alpha)}{q^3} \left[-\frac{g''(q)}{g(q)^2} q + \frac{2g'(q)^2}{g(q)^3} + 2\frac{g(q)^4}{g'(q)} \right].$$

- As soon as g is increasing and concave, φ must be convex.

- Theorem 1:

- ▶ The distortion risk measure $\rho_g(X) = \int_0^{+\infty} g(1 - F_X(x))dx$ is subadditive if, and only if, g is a concave distortion function, without strictly convex sections.

- Theorem 2:

- ▶ A Haezendonck-Goovaerts risk measure risk measure, with φ derived from a concave distortion function g is subadditive.

The generalized Haezendonck-Goovaerts risk measure risk measure

- Residual risk:

- ▶ The solution of

$$E \left[\frac{\left(F_X^{-1}(U) - F_X^{-1}(\alpha) \right)_+}{\rho_l - F_X^{-1}(\alpha)} \right] = 1 - \alpha. \quad (12)$$

- ▶ is $\rho_l \left(X, F_X^{-1}(\alpha) \right)$:

$$\rho_l \left(X, F_X^{-1}(\alpha) \right) = \rho_l - F_X^{-1}(\alpha) = \frac{1}{1 - \alpha} E U \left[\left(F_X^{-1}(U) - F_X^{-1}(\alpha) \right)_+ \right].$$

- A generalization:

- ▶ $g(x) = \min \left\{ \frac{x}{1 - \alpha}, 1 \right\}$.
- ▶ Corresponding distortion risk measure:

$$\rho_g \left(X, F_X^{-1}(\alpha) \right) = \int_0^1 F_X^{-1}(y) g'(1 - y) dy.$$

The generalized Haezendonck-Goovaerts risk measure risk measure (cont'd)

- Consider the following relation:

$$\mathbb{E} \left[\frac{(F_X^{-1}(U) - F_X^{-1}(\alpha))_+ g'(1-U)}{\rho_{l,g} - F_X^{-1}(\alpha)} \right] = 1 - \alpha. \quad (13)$$

- Solving for ρ gives:

$$\begin{aligned} \rho_{l,g} - F_X^{-1}(\alpha) &= \frac{1}{1-\alpha} \mathbb{E} \left[(F_X^{-1}(U) - F_X^{-1}(\alpha))_+ g'(1-U) \right] \\ &= \frac{1}{1-\alpha} \int_{\alpha}^1 (F_X^{-1}(u) - F_X^{-1}(\alpha)) g'(1-u) du \\ &= \frac{1}{1-\alpha} \int_{\alpha}^1 F_X^{-1}(u) g'(1-u) du \\ &\quad - \frac{1}{1-\alpha} \int_{\alpha}^1 F_X^{-1}(\alpha) g'(1-u) du. \end{aligned}$$

The generalized Haezendonck-Goovaerts risk measure risk measure (cont'd)

- Solution:

$$\begin{aligned}\rho_{l,g} &= F_X^{-1}(\alpha) + \frac{1}{1-\alpha} \int_0^1 F_X^{-1}(u) g'(1-u) du - \frac{F_X^{-1}(\alpha)}{1-\alpha} \\ &= F_X^{-1}(\alpha) + \frac{1}{1-\alpha} \left(\rho_g(X, F_X^{-1}(\alpha)) - F_X^{-1}(\alpha) \right).\end{aligned}\quad (14)$$

- Conclusion:

- ▶ A distortion function can be used to generate risk measures, which are solutions of equation

$$E \left[\frac{\left(F_X^{-1}(U) - F_X^{-1}(\alpha) \right)_+ g'(1-U)}{\rho_{l,g} - F_X^{-1}(\alpha)} \right] = 1 - \alpha. \quad (15)$$

- ▶ This solution can be linked with the solution of equation

$$E \left[\frac{\left(F_X^{-1}(U) - F_X^{-1}(\alpha) \right)_+}{\rho_{l,g} - F_X^{-1}(\alpha)} \right] = 1 - \alpha.$$

The generalized Haezendonck-Goovaerts risk measure risk measure (cont'd)

- Definition:

- ▶ Let φ be a strictly increasing function with $\varphi(0) = 0$, $\varphi(1) = 1$, $\varphi(+\infty) = +\infty$ and let $\alpha \in (0, 1)$. Given a distortion function g , which is an increasing function, satisfying $g(0) = 0$ and $g(1) = 1$, the generalized Haezendonck-Goovaerts risk measure is denoted by $\rho_{\varphi,g}(X, t)$ and is the solution of:

$$1 - \alpha = E \left[\varphi \left(\frac{\left(F_X^{-1}(U) - t \right)_+}{\rho_{\varphi,g}(X, t) - t} \right) g'(1 - U) \right].$$

- Theorem:

- ▶ Let φ be a continuous and strictly increasing function in \mathbb{R}^+ and g a valid distortion function. A necessary and sufficient condition that $\rho_{\varphi,g}(X, t)$ is comparable larger than $\rho_{l,g}(X, t)$ (l means a linear φ) is that φ is a convex function.

Proof

- Use $\rho_{\varphi,g} = \rho_{\varphi,g}(X, t)$ and $\rho_{l,g} = \rho_{l,g}(X, t)$.
- Assume: $\rho_{\varphi,g} > \rho_{l,g}$

$$1 - \alpha = \varphi(1 - \alpha) = \mathbb{E}_U \left[\varphi \left(\frac{(F_X^{-1}(U) - F_X^{-1}(\alpha))_+ \rho_{l,g}}{\rho_{\varphi,g}} \right) g'(1 - U) \right].$$

- Since $\frac{\rho_{l,g}}{\rho_{\varphi,g}} < 1$ we get

$$1 - \alpha < \mathbb{E}_U \left[\varphi \left(\frac{(F_X^{-1}(U) - F_X^{-1}(\alpha))_+}{\rho_{l,g}} \right) g'(1 - U) \right].$$

Proof (cont'd)

- On the other hand one has (using the definition of $\rho_{l,g}$) that

$$\begin{aligned} 1 - \alpha &= \mathbb{E}_U \left[\varphi \left(\frac{(F_X^{-1}(U) - F_X^{-1}(\alpha))_+}{\rho_{l,g}} \right) g'(1 - U) \right] \\ &= \varphi \left(\mathbb{E}_U \left[\frac{(F_X^{-1}(U) - F_X^{-1}(\alpha))_+}{\rho_{l,g}} g'(1 - U) \right] \right). \end{aligned}$$

- Hence $\mathbb{E}[\varphi(Z)] > \varphi(\mathbb{E}[Z])$ and φ is convex.
- In case φ is convex one immediately sees that $\rho_{\varphi,g} > \rho_{l,g}$, as for every l one gets the inequality

$$\begin{aligned} &\mathbb{E} \left[\varphi \left(\frac{(F_X^{-1}(U) - F_X^{-1}(\alpha))_+}{\rho} \right) g'(1 - U) \right] \\ &> \varphi \left(\mathbb{E} \left[\frac{(F_X^{-1}(U) - F_X^{-1}(\alpha))_+}{\rho} g'(1 - U) \right] \right). \end{aligned}$$

Subadditivity, comonotonicity and monotonicity

- Consider the following axioms:

- ▶ $X <_{cx} Y \implies \rho_\varphi(X, t) \geq \rho_\varphi(Y, t)$ implies that φ is convex.
- ▶ In case φ is determined such that for a two point risk the resulting risk measure is concave, φ is convex.
- ▶ Comparability of a Haezendonck-Goovaerts risk measure risk measure and the conditional tail expectations implies that φ has to be convex.

- In the framework of the Haezendonck risk measure, they are "equivalent" in the sense that they imply convexity of the function φ :

Subadditivity, comonotonicity and monotonicity (cont'd)

- Theorem:

- ▶ Let us consider a discrete cumulative distribution function after the inclusion of a shift t at the origin. In fact, one considers a discrete random variable $(X - t)_+$ with

$$\Pr((X - t)_+ = a_i) = q_i, \quad i = 1, 2, \dots, n.$$

Suppose the function φ is determined by a distortion function g such that

$$\frac{1 - \alpha}{q} = \varphi\left(\frac{1}{g(q)}\right),$$

for $q < 1 - \alpha$ and $\Pr((X - t)_+ = 0) = 1 - \sum_{i=1}^n q_i$. Then, the Haezendonck-Goovaerts risk measure with a convex function φ provides an upper bound for the distortion risk measure.

Proof

- Define the Bernoulli random variable $X_{(a_i - a_{i-1})q_i}$ as follows:

$$\Pr \left(X_{(a_i - a_{i-1})q_i} = a_i - a_{i-1} \right) = q_i,$$
$$\Pr \left(X_{(a_i - a_{i-1})q_i} = 0 \right) = 1 - q_i.$$

- In this case,

$$F_X^{-1}(U) \stackrel{d}{=} \sum_{i=1}^n X_{(a_i - a_{i-1})q_i}$$
$$\stackrel{d}{=} \sum_{i=1}^n F_{X_{(a_i - a_{i-1})q_i}}^{-1}(U).$$

- The distortion risk measure with distortion function g :

$$\rho_l((X - b)_+) = \sum_{i=1}^n (a_i - a_{i-1}) g(q_i),$$

Proof (cont'd)

- consequently:

$$\begin{aligned} & E \left[\varphi \left(\frac{\sum_{i=1}^n F_{X_{(a_i - a_{i-1})q_i}}^{-1}(U)}{\sum_{i=1}^n (a_i - a_{i-1}) g(q_i)} \right) \right] \\ & \geq \sum_{i=1}^n \frac{(a_i - a_{i-1}) g(q_i)}{\sum_{j=1}^n (a_j - a_{j-1}) g(q_j)} E \left[\varphi \left(\frac{F_{X_{(a_i - a_{i-1})q_i}}^{-1}(U)}{(a_i - a_{i-1}) g(q_i)} \right) \right] \\ & = \sum_{i=1}^n \frac{(a_i - a_{i-1}) g(q_i)}{\sum_{j=1}^n (a_j - a_{j-1}) g(q_j)} q_i \varphi \left(\frac{1}{g(q_i)} \right) \\ & = 1 - \alpha. \end{aligned}$$

- Hence

$$E \left[\varphi \left(\frac{\sum_{i=1}^n F_{X_{(a_i - a_{i-1})q_i}}^{-1}(U)}{\rho_\varphi((X - t)_+)} \right) \right] = 1 - \alpha,$$

entails that $\rho_\varphi((X - t)_+)$ is larger than the distortion risk measure

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